## HW8

## April 10, 2024

**Exercise 1** (KS, Ex 3.5.18) Let  $B = (B_t)_{0 \le t \le 1}$  be a Brownian motion. Define

 $T = \inf\{0 \le t \le 1: t + B_t^2 = 1\}$ 

and

$$X_t = -\frac{2}{(1-t)^2} B_t \mathbb{1}_{\{t \le T, t < 1\}}, \quad 0 \le t < 1.$$

1. Show that  $\mathsf{P}(T < 1) = 1$ , and hence  $\int_0^1 X_t^2 dt < \infty$  a.s. proof . By definition of T, we know  $0 \leq T(w) \leq 1$  for wED. Suppose those exists a subset  $A \subseteq \Omega$  with P(A) > 0such that T(w)=1 for wEA, then by continuity of B(., w) we have for wEA. 1> Two + BT (w) = 1+B7 (w), which means that B1(w) = 0 for wEA. It contradicts the fact that B1 follows the distribution Normal (0,1) (which is absolutely continuous with respect to the Lebesgue measure on IR). Therefore P(T<1)=1. Then for a.s. w, Twos<1, which implies that =  $\left| X_{t} \right|^{2} = \left[ -\frac{2}{(1-t)^{2}} B_{t} 1_{\{t \leq T, t < i\}} \right]^{2} = \frac{4}{(1-t)^{4}} B_{t}^{2}(w) 1_{\{t \leq T, t < i\}} =$  $\frac{4}{(1-t)^4}B_t^{2(w)} I_{qt} \leq T_y^{(t,w)} \text{ is bounded in } t \in [0,1) \text{ for a.s. } w \in \Omega,$ which means that  $\int_0^1 X_t^2 dt < \infty$  for a.s.  $w \in \Omega$ .

2. Apply Itô's formula to the process  $(1-t)^{-2}B_t^2$  to conclude that

$$\int_0^1 X_t \, dB_t - \frac{1}{2} \int_0^1 X_t^2 \, dt = -1 - 2 \int_0^T \left[ \frac{1}{(1-t)^4} - \frac{1}{(1-t)^3} \right] B_t^2 \, dt \le -1.$$

3. Show that the exponential super-martingale  $Z_t(X)$ ,  $0 \le t \le 1$  is not a martingale; however, for each  $n \ge 1$  and  $\sigma_n = 1 - (1/\sqrt{n})$ ,  $Z_{t \land \sigma_n}(X)$ ,  $0 \le t \le 1$  is a martingale.

Proof. 
$$Z_{t}(X) = \exp\left(\int_{0}^{t} X_{s} dB_{s} - \frac{1}{2} \int_{0}^{t} X_{s}^{2} dS\right)$$
. Obviously we have  
that  $Z_{0}(X) = 1$ , which implies that  $EZ_{1}(X) = 1$ . But by 2)  
we obtain that  $Z_{1}(X) \leq \exp(-1)$ , which means that  
 $EZ_{0}(X) = 1 > \exp(-1) \gg EZ_{1}(X)$ , which implies  
that  $(Z_{t}(X))_{t \in [0,1)}$  is not a martingale.

To show that 
$$(\mathbb{Z}_{t \wedge Gn})_{t \in [0,1]}$$
 is a martingale for each  $n \ge 1$ , by Novikov's contenion  
it suffices to show  $\mathbb{E}[\exp(\pm \int_{0}^{Gn} \chi_{t}^{2} dt)] < \infty$ . We have  
 $\mathbb{E}[\exp(\pm \int_{0}^{Gn} \chi_{t}^{2} dt)] = \mathbb{E}[\exp(\frac{1}{2} \int_{0}^{Gn} \frac{4}{(1-t)^{4}} B_{t}^{2} 1_{\{t \le T\}} dt)]$   
 $\leq \mathbb{E}[\exp(2 \int_{0}^{Gn} \frac{1}{(1-t)^{3}} dt)] = \exp(2 \int_{0}^{Gn} \frac{1}{(1-t)^{3}} dt) < \infty$ .  
The proof is therefore done.

**Exercise 2** (Le Gall, Ex 5.28) Let B be a Brownian motion started from 1. Fix  $\varepsilon \in (0, 1)$  and set  $T_{\varepsilon} = \inf\{t \ge 0 : B_t = \varepsilon\}$ . Also let  $\lambda > 0$  and  $\alpha \in \mathbb{R} \setminus \{0\}$ .

1. Show that  $Z_t = (B_{t \wedge T_{\varepsilon}})^{\alpha}$  is a semi-martingale and give its canonical decomposition as the sum of a c.l.m. and a finite variation process.

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Proof. For 
$$t \ge 0$$
,  $B_{t} \wedge T_{\varepsilon} = \int_{0}^{t \wedge T_{\varepsilon}} 1 dB_{s} = \int_{0}^{t} 1_{1} g \le T_{\varepsilon} g dB_{s}$   
Then by Itô's formula,  
 $Z_{t} = (B_{t} \wedge T_{\varepsilon})^{\alpha} = 1 + \int_{0}^{t} \alpha (B_{s} \wedge T_{\varepsilon})^{\alpha-1} 1_{\{s \le T_{\varepsilon}\}} dB_{s}$   
 $+ \int_{0}^{t} \frac{1}{2} \alpha (\alpha - 1) (B_{s} \wedge T_{\varepsilon})^{\alpha-1} 1_{\{s \le T_{\varepsilon}\}} dS$ . Set for  $t \ge 0$  that  
 $M_{t} := \int_{0}^{t} \alpha (B_{s} \wedge T_{\varepsilon})^{\alpha-1} 1_{\{s \le T_{\varepsilon}\}} dB_{s}$ , which is a c.l.m. and  
 $A_{t} := \int_{0}^{t} \frac{1}{2} \alpha (\alpha - 1) (B_{s} \wedge T_{\varepsilon})^{\alpha-2} 1_{\{s \le T_{\varepsilon}\}} dS$ , which is a f.v. process.

2. Show that the process

$$Z_t = (B_{t \wedge T_{\varepsilon}})^{\alpha} \exp\left(-\lambda \int_0^{t \wedge T_{\varepsilon}} \frac{ds}{B_s^2}\right)$$

is a c.l.m. if  $\alpha$  and  $\lambda$  satisfy a polynomial equation to be determined.

Proof For 
$$t \ge 0$$
,  $Z_t = (B_{t \land T_E})^{\alpha} \exp\left(-\lambda \int_0^t \frac{1}{B_s^2} ds\right)$   

$$= 1 + \int_0^t (B_{s \land T_E})^{\alpha} \exp\left(-\lambda \int_0^s \frac{1}{B_t^2} d\tau\right) \cdot (-\lambda \cdot 1_{f \le T_E y} \cdot \frac{1}{B_s^2}) ds$$

$$+ \int_0^t \exp\left(\lambda \int_0^s \frac{1}{B_t^2} d\tau\right) \cdot \alpha (B_{s \land T_E})^{\alpha-1} 1_{f \le T_E y} dB_s$$

$$+ \int_0^t \exp\left(\lambda \int_0^s \frac{1}{B_t^2} (\tau \le T_E y) d\tau\right) \cdot \alpha (B_{s \land T_E})^{\alpha-1} 1_{f \le T_E y} dB_s$$

$$+ \int_0^t \exp\left(\lambda \int_0^s \frac{1}{B_t^2} (\tau \le T_E y) d\tau\right) \cdot \alpha (B_{s \land T_E})^{\alpha-1} 1_{f \le T_E y} dB_s$$

$$T_0 \text{ make } (Z_t)_{t \ge 0} a \quad c \cdot I \cdot m. \text{, it's equivalent to have : For  $s \ge 0$ , (B_{s \land T_E})^{\alpha} (-\lambda) 1_{f \le T_E y} \cdot \frac{1}{B_{s \land T_E}} + \frac{1}{2} \alpha (\alpha-1) (B_{s \land T_E})^{\alpha-2} 1_{f \le T_E y} = 0$$
, which means that  $\lambda = \frac{1}{2} \alpha (\alpha-1)$ .

3. Compute

$$\mathsf{E}\Big[\exp\big(-\lambda\int_0^{T_\varepsilon}\frac{ds}{B_s^2}\big)\Big].$$

Proof. Let Z<sub>t</sub> be the same as in 2). And let 
$$\alpha \in \operatorname{IR}(\operatorname{for})$$
  
such that  $\lambda = \frac{1}{2}\alpha(\alpha - 1)$ . (It always exists such a since  $\lambda > 0$ ). There exists  
a localizing sequence  $(S_n)_{n \ge 1}$  such that for each  $n \ge 1$ ,  $(\mathbb{Z} + A S_n)_{\substack{t \ge 0 \\ n \ge 1}}$   
is a uniformly integrable martingale. Since T<sub>e</sub> is a stopping-time,  
by 0.5.T. We have  $\mathbf{1} = \mathbb{E}\mathbb{Z}_0 = \mathbb{E}\mathbb{Z}_{0AS_n} = \mathbb{E}\mathbb{Z}_{TeAS_n} =$   
 $\mathbb{E}[(B_{TeAS_n})^{\alpha} \exp(-\lambda \int_0^{TeAS_n} \frac{1}{B_s^2} ds)]$ . Note that  
 $\lim_{n \to +\infty} (B_{TeAS_n})^{\alpha} \exp(-\lambda \int_0^{TeAS_n} \frac{1}{B_s^2} ds) = B_{Te}^{\alpha} \exp(-\lambda \int_0^{Te} \frac{1}{B_s^2} ds)$ ,  
and for each  $n \ge 1$ ,  $|(B_{TeAS_n})^{\alpha} \exp(-\lambda \int_0^{TeAS_n} \frac{ds}{B_e^2})| \le (B_{TeAS_n})^{\alpha}$ .  
By the equation  $\lambda = \frac{1}{2}\alpha(\alpha - 1)$ , we have  $\alpha = \frac{1 \pm \sqrt{1 + 8\lambda}}{2}$ .  
If we take  $\alpha = \frac{1 - \sqrt{1 + 8\lambda}}{2} < 0$ , then we have  $(B_{TeAS_n})^{\alpha} \le B_{Te}^{\alpha} = \varepsilon^{\alpha}$  and by D.C.T. we have that  $1 = \mathbb{E}[B_{Te}^{\alpha} \exp(-\lambda \int_{0}^{Te} \frac{ds}{B_s^2})]$   
 $= \varepsilon^{\alpha} \mathbb{E}[\exp(-\lambda \int_{0}^{Te} \frac{ds}{B_s^2}] = \varepsilon^{-\alpha}$ , where  $\alpha = \frac{1 - \sqrt{1 + 8\lambda}}{2}$ .

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